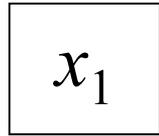
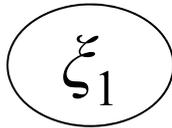


Notation and preliminaries

Psychology 588: Covariance structure and factor models



observed variables



latent variables

δ_1

error terms (unenclosed)



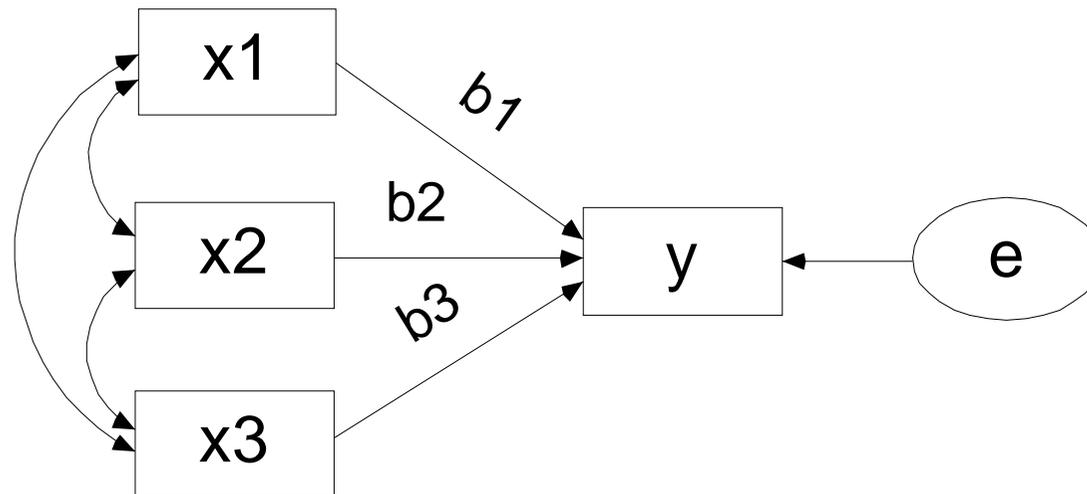
causal path or factor loading/weight



covariance or unconstrained (nonzero) relationship
between exogenous variables

$$y = b_1x_1 + b_2x_2 + b_3x_3 + e$$

$$= \mathbf{b}'\mathbf{x} + e$$

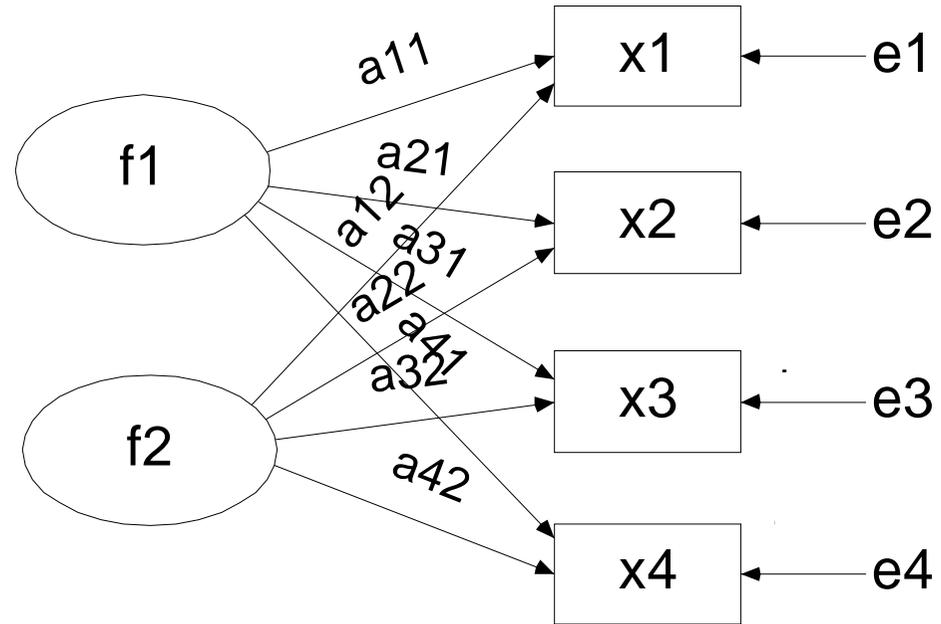


- How could the intercept term enter into the diagram?

Principal component model (rank-reduced)

$$\mathbf{x} = \mathbf{A}\mathbf{f} + \mathbf{e}$$

$$\text{cov}(\mathbf{x}) = \mathbf{A}\mathbf{A}' + \mathbf{\Theta}$$



- 2 orthogonal components fitted to 4 manifest variables
- Rotationally indeterminate if no component loadings are constrained (e.g., to zero)

-
- Like the standard matrix notation, the LISREL notation uses
 - Lower italic: a scalar (as a variable or a parameter; *cf.* X)
 - Lower bold: vector
 - Upper bold: matrix
 - Uses Roman letters only for observed (or manifest, indicator) variables (e.g., \mathbf{x} , \mathbf{y}) and Greek letters for all others (i.e., latent variables and model parameters; ξ , Γ)
 - Distinguishes between exogenous (independent) vs. endogenous (dependent) variables parts
 - We will consider later an alternative notation and model representation --- the reticular action model (RAM, chap. 9)

Symbol	size	Definition
Variables		
ξ	$n \times 1$	latent exogenous variables
η	$m \times 1$	latent endogenous variables
ζ	$m \times 1$	specification error terms
Coefficients		
Γ	$m \times n$	coefficient matrix for ξ
\mathbf{B}	$m \times m$	coefficient matrix for η
Covariance matrix		
Φ	$n \times n$	covariance matrix of ξ , $E(\xi\xi')$
Ψ	$m \times m$	covariance matrix of ζ , $E(\zeta\zeta')$

Symbol	size	Definition
variables		
\mathbf{x}	$q \times 1$	observed indicators of ξ
\mathbf{y}	$p \times 1$	observed indicators of η
δ	$q \times 1$	measurement errors for \mathbf{x}
ε	$p \times 1$	measurement errors for \mathbf{y}
coefficients		
Λ_x	$q \times n$	loadings relating \mathbf{x} to ξ
Λ_y	$p \times m$	loadings relating \mathbf{y} to η
covariance matrix		
Θ_δ	$q \times q$	covariance matrix of δ , $E(\delta\delta')$
Θ_ε	$p \times p$	covariance matrix of ε , $E(\varepsilon\varepsilon')$

$$\boldsymbol{\eta} = \mathbf{B}\boldsymbol{\eta} + \boldsymbol{\Gamma}\boldsymbol{\xi} + \boldsymbol{\zeta} \quad \text{or} \quad \boldsymbol{\eta} = (\mathbf{I} - \mathbf{B})^{-1} (\boldsymbol{\Gamma}\boldsymbol{\xi} + \boldsymbol{\zeta})$$

$$\begin{bmatrix} \eta_1 \\ \vdots \\ \eta_m \end{bmatrix} = \begin{bmatrix} 0 & \beta_{12} & \cdots & \beta_{1m} \\ \beta_{21} & \ddots & & \vdots \\ \vdots & & \ddots & \beta_{m-1,m} \\ \beta_{m1} & \cdots & \beta_{m,m-1} & 0 \end{bmatrix} \begin{bmatrix} \eta_1 \\ \vdots \\ \eta_m \end{bmatrix} + \begin{bmatrix} \gamma_{11} & \cdots & \gamma_{1n} \\ \vdots & \ddots & \vdots \\ \gamma_{m1} & \cdots & \gamma_{mn} \end{bmatrix} \begin{bmatrix} \xi_1 \\ \vdots \\ \xi_n \end{bmatrix} + \begin{bmatrix} \zeta_1 \\ \vdots \\ \zeta_m \end{bmatrix}$$

$$\eta_i = \left(\cdots + \beta_{i(i-2)}\eta_{i-2} + \beta_{i(i-1)}\eta_{i-1} \right) + \left(\gamma_{i1}\xi_1 + \gamma_{i2}\xi_2 + \cdots + \gamma_{in}\xi_n \right) + \zeta_i$$

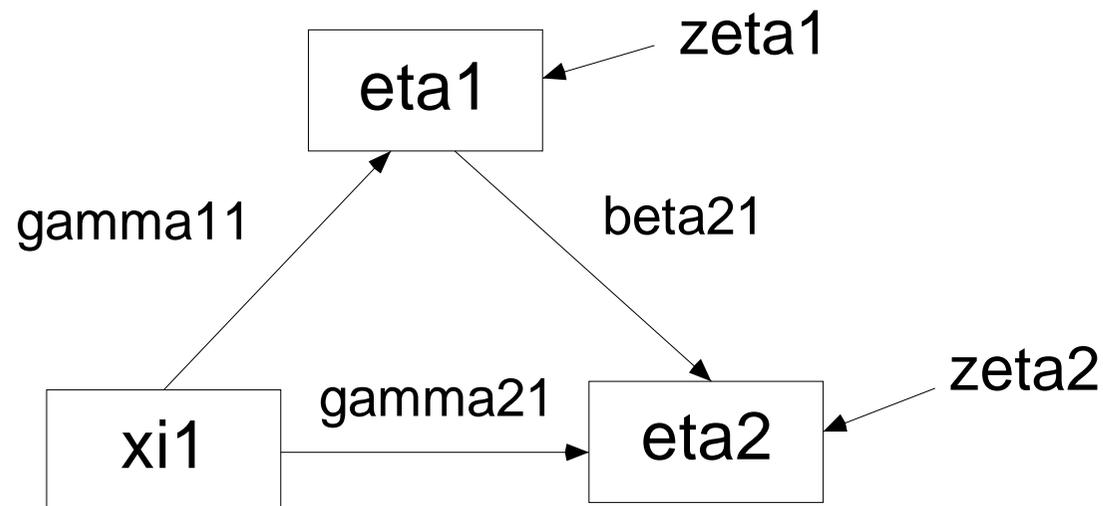
- Equations are linear both in the variables (η , ξ) and in the parameters (β , γ); same for equations for manifest variables

- $E(\boldsymbol{\eta}) = \mathbf{0}, \quad E(\boldsymbol{\xi}) = \mathbf{0}, \quad E(\boldsymbol{\zeta}) = \mathbf{0}$
- $E(\boldsymbol{\zeta}\boldsymbol{\zeta}') = \mathbf{0}_{m \times n}$
- $(\mathbf{I} - \mathbf{B})$ is nonsingular so that $(\mathbf{I} - \mathbf{B})^{-1}$ exists
- ζ_i has a homogeneous variance for all subjects, i.e.,
$$E(\zeta_{ik}^2) = \text{var}(\zeta_i) \quad \text{for } k = 1, \dots, N, \quad i = 1, \dots, m$$

and independent, i.e., $\text{cov}(\zeta_{ik}, \zeta_{il}) = 0$ for all $k \neq l$;
otherwise, multilevel structure

$$\boldsymbol{\eta} = \mathbf{B}\boldsymbol{\eta} + \boldsymbol{\Gamma}\boldsymbol{\xi} + \boldsymbol{\zeta}$$

$$\begin{bmatrix} \eta_1 \\ \eta_2 \end{bmatrix} = \begin{bmatrix} 0 & 0 \\ \beta_{21} & 0 \end{bmatrix} \begin{bmatrix} \eta_1 \\ \eta_2 \end{bmatrix} + \begin{bmatrix} \gamma_{11} \\ \gamma_{21} \end{bmatrix} \xi_1 + \begin{bmatrix} \zeta_1 \\ \zeta_2 \end{bmatrix}$$



$$\mathbf{x} = \Lambda_x \boldsymbol{\xi} + \boldsymbol{\delta}$$

$$\mathbf{y} = \Lambda_y \boldsymbol{\eta} + \boldsymbol{\varepsilon}$$

- Often, a uni-factorial structure is imposed on Λ_x, Λ_y such as:

$$\begin{bmatrix} x_1 \\ \vdots \\ x_q \end{bmatrix} = \begin{bmatrix} \lambda_{11} & \cdots & \lambda_{1n} \\ \vdots & \ddots & \vdots \\ \lambda_{q1} & \cdots & \lambda_{qn} \end{bmatrix} \begin{bmatrix} \xi_1 \\ \vdots \\ \xi_n \end{bmatrix} + \begin{bmatrix} \delta_1 \\ \vdots \\ \delta_q \end{bmatrix}$$

$$x_i = \lambda_{i1} \xi_1 + \lambda_{i2} \xi_2 + \cdots + \lambda_{in} \xi_n + \delta_i$$

$$\Lambda_x = \begin{bmatrix} \lambda_{11} & 0 & 0 \\ \lambda_{21} & 0 & 0 \\ \lambda_{31} & 0 & 0 \\ 0 & \lambda_{42} & 0 \\ 0 & \lambda_{52} & 0 \\ 0 & \lambda_{62} & 0 \\ 0 & 0 & \lambda_{73} \\ 0 & 0 & \lambda_{83} \\ 0 & 0 & \lambda_{93} \end{bmatrix}$$

- All latent variables expected to be zero --- a natural consequence when fitting mean-centered manifest variables

$$E(\boldsymbol{\eta}) = \mathbf{0}, \quad E(\boldsymbol{\xi}) = \mathbf{0}, \quad E(\boldsymbol{\varepsilon}) = \mathbf{0}, \quad E(\boldsymbol{\delta}) = \mathbf{0}$$

But not so natural when multiple groups or hierarchically nested (multilevel) data are considered

- Observed variables are correlated only through the modeled latent variables and parameters (latent path coefficients and loadings), with simplifying conditions of

$$E(\boldsymbol{\varepsilon}\boldsymbol{\eta}') = \mathbf{0}, \quad E(\boldsymbol{\varepsilon}\boldsymbol{\xi}') = \mathbf{0}, \quad E(\boldsymbol{\varepsilon}\boldsymbol{\zeta}') = \mathbf{0}, \quad E(\boldsymbol{\varepsilon}\boldsymbol{\delta}') = \mathbf{0}$$

$$E(\boldsymbol{\delta}\boldsymbol{\xi}') = \mathbf{0}, \quad E(\boldsymbol{\delta}\boldsymbol{\eta}') = \mathbf{0}, \quad E(\boldsymbol{\delta}\boldsymbol{\zeta}') = \mathbf{0}$$

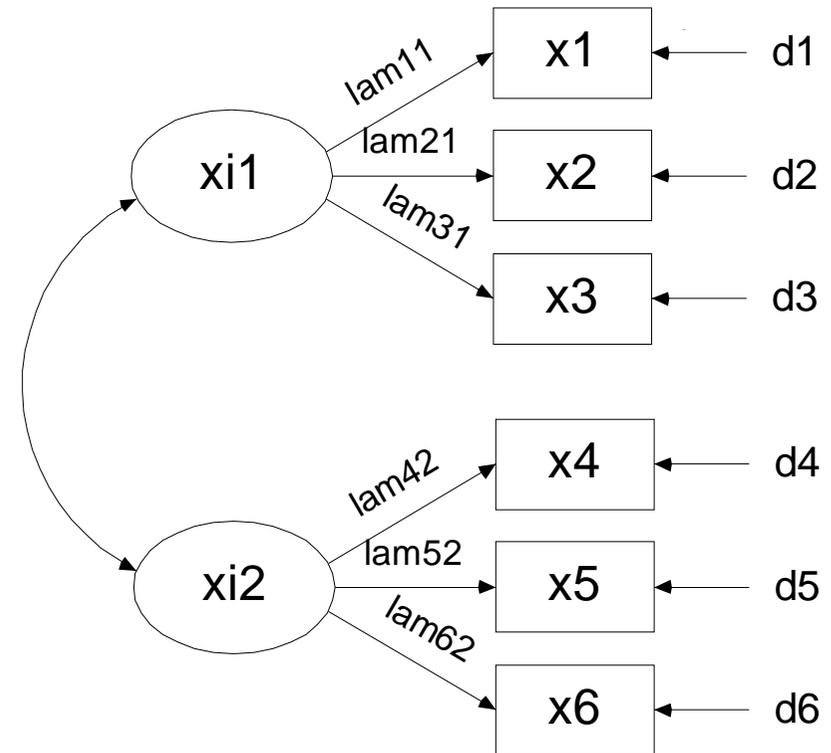
- Typically, Θ_δ and Θ_ε are considered to be diagonal matrices, i.e., measurement errors of the indicators are uncorrelated; however, there may be justifiable cases for non-diagonal Θ_δ and Θ_ε or even $E(\varepsilon\delta') \neq \mathbf{0}$, for some selective entries

e.g., Fig. 2.6 (p. 37) of industrialization and democracy model,

- Like ζ , δ and ε are assumed to be homoscedestic and independent (i.e., iid)
- We will consider later distributional assumptions for manifest variables \mathbf{x} and \mathbf{y}

$$\mathbf{x} = \Lambda_x \boldsymbol{\xi} + \boldsymbol{\delta}$$

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \\ x_5 \\ x_6 \end{bmatrix} = \begin{bmatrix} \lambda_{11} & 0 \\ \lambda_{21} & 0 \\ \lambda_{31} & 0 \\ 0 & \lambda_{42} \\ 0 & \lambda_{52} \\ 0 & \lambda_{62} \end{bmatrix} \begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} + \begin{bmatrix} \delta_1 \\ \delta_2 \\ \delta_3 \\ \delta_4 \\ \delta_5 \\ \delta_6 \end{bmatrix}$$



- Where is the double-headed arrow defined in the equation?

- Following rules may be useful:

$$\text{cov}(c, X) = 0$$

$$\text{cov}(cX_1, X_2) = c \cdot \text{cov}(X_1, X_2)$$

$$\text{cov}(X_1 + X_2, X_3) = \text{cov}(X_1, X_3) + \text{cov}(X_2, X_3)$$

- But the standard algebra will do; suppose, e.g., $x_1 = \lambda_1 \xi_1 + \delta_1$,
 $x_2 = \lambda_2 \xi_1 + \delta_2$:

$$\begin{aligned} \text{cov}(x_1, x_2) &= E(x_1 x_2) = E(\lambda_1 \xi_1 + \delta_1)(\lambda_2 \xi_1 + \delta_2) \\ &= E(\lambda_1 \lambda_2 \xi_1 \xi_1 + \lambda_1 \delta_2 \xi_1 + \lambda_2 \delta_1 \xi_1 + \delta_1 \delta_2) \\ &= \lambda_1 \lambda_2 E(\xi_1 \xi_1) = \lambda_1 \lambda_2 \phi_{11} \end{aligned}$$

- Matrix algebra useful for operation of covariance matrices:

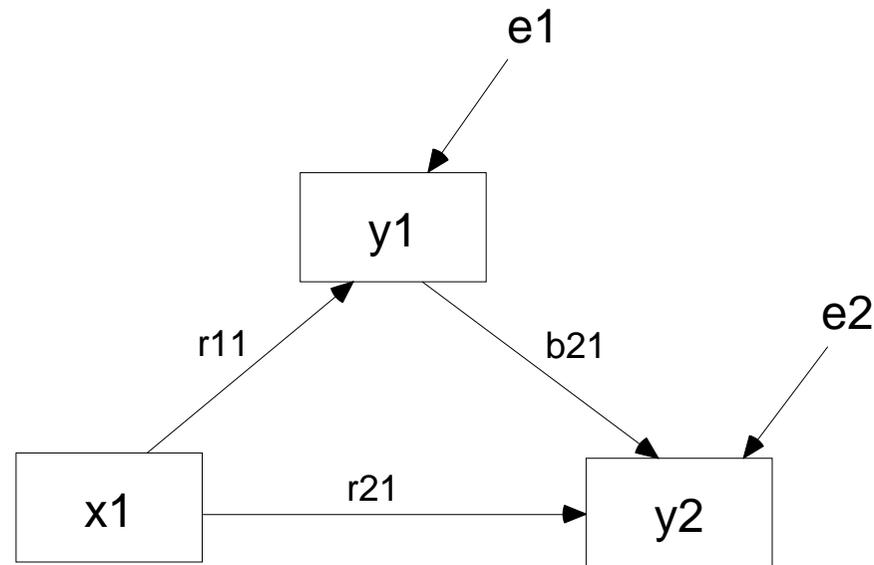
$$\mathbf{x} = \mathbf{\Lambda}_x \boldsymbol{\xi} + \boldsymbol{\delta},$$

$$\text{cov}(\mathbf{x}, \mathbf{x}) = E(\mathbf{x}\mathbf{x}') = E(\mathbf{\Lambda}_x \boldsymbol{\xi} + \boldsymbol{\delta})(\boldsymbol{\xi}'\mathbf{\Lambda}_x' + \boldsymbol{\delta}')$$

$$= E(\mathbf{\Lambda}_x \boldsymbol{\xi} \boldsymbol{\xi}' \mathbf{\Lambda}_x' + \mathbf{\Lambda}_x \boldsymbol{\xi} \boldsymbol{\delta}' + \boldsymbol{\delta} \boldsymbol{\xi}' \mathbf{\Lambda}_x' + \boldsymbol{\delta} \boldsymbol{\delta}')$$

$$= \mathbf{\Lambda}_x E(\boldsymbol{\xi} \boldsymbol{\xi}') \mathbf{\Lambda}_x' + E(\boldsymbol{\delta} \boldsymbol{\delta}') = \mathbf{\Lambda}_x \boldsymbol{\Phi} \mathbf{\Lambda}_x' + \boldsymbol{\Theta}_\delta$$

- Direct effect: unmediated expected change on a variable due to another --- a path coefficient
- Indirect effect: all other possible influences from a variable to another, other than its direct effect --- product of all involved mediating path coefficients for each indirect effect
- Total effect = direct effect + all indirect effects
- Recursive vs. non-recursive models
- convergent vs. divergent series of indirect effects

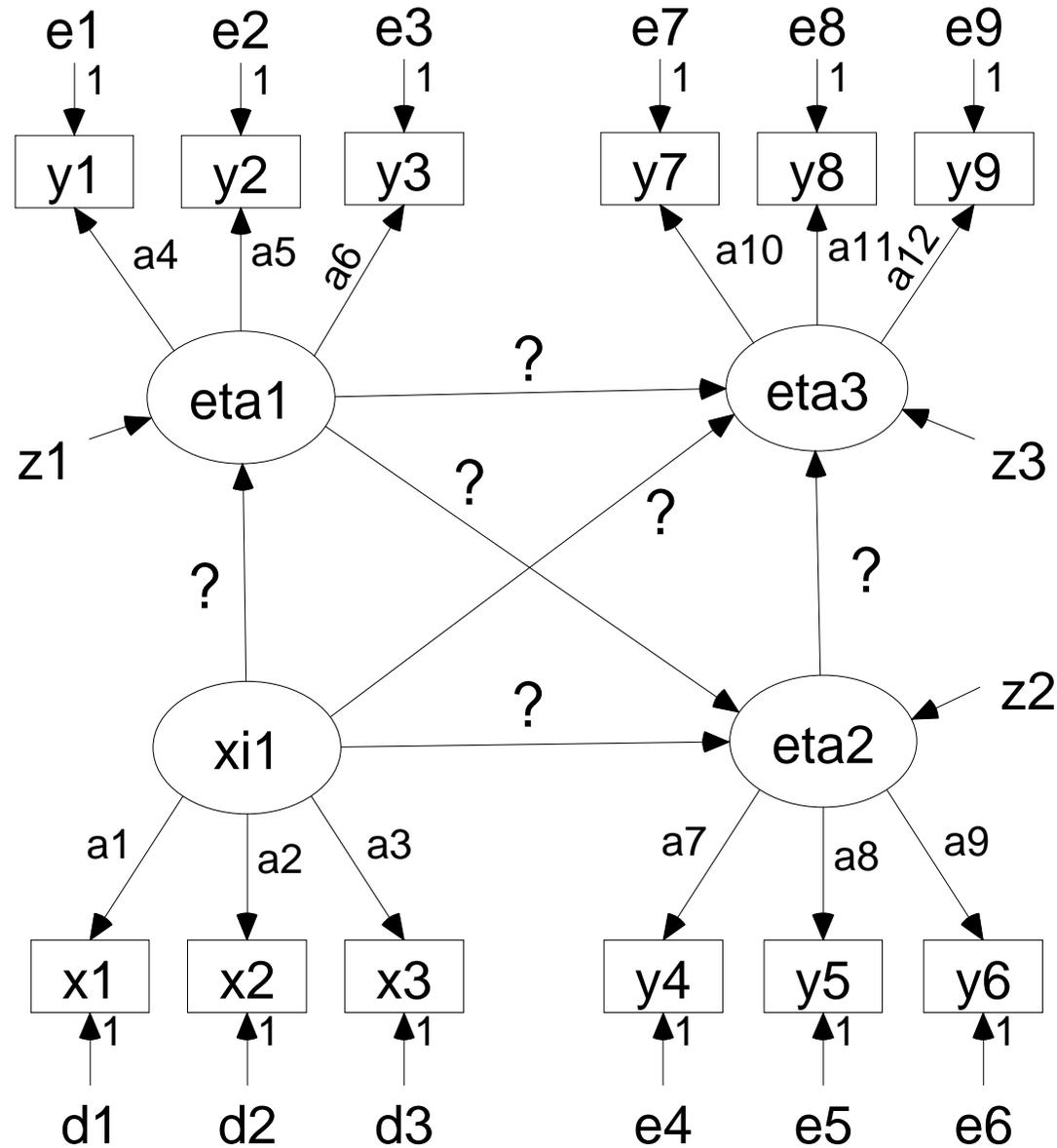


- Direct effect of x_1 on $y_2 = \gamma_{21}$

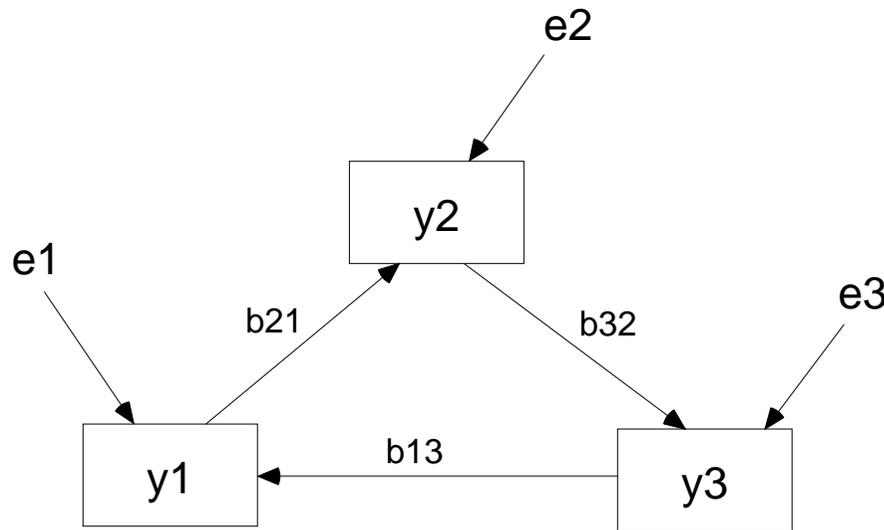
Indirect effect = $\gamma_{11}\beta_{21}$ --- expected change on y_2 due to 1 unit change on x_1

Total effect = $\gamma_{21} + \gamma_{11}\beta_{21}$

Effect decomposition example 2



- non-recursive since all 3 variables indirectly influence itself



$$\mathbf{B} = \begin{bmatrix} 0 & 0 & \beta_{13} \\ \beta_{21} & 0 & 0 \\ 0 & \beta_{32} & 0 \end{bmatrix}$$

- Indirect effect of y_1 on y_3 :

$$\begin{aligned}
 & \beta_{21}\beta_{32} + \beta_{21}\beta_{32}(\beta_{13}\beta_{21}\beta_{32}) + \beta_{21}\beta_{32}(\beta_{13}\beta_{21}\beta_{32})^2 + \dots \\
 & = \beta_{21}\beta_{32} \sum_{k=0}^{\infty} (\beta_{13}\beta_{21}\beta_{32})^k
 \end{aligned}$$